Do People Accept Different Cultures?

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Abstract
We present a model of the ethnic preferences of a minority group of immigrants and a majority group of natives for different cultures. We show that ethnic preferences change when there is an increase in minority populations or when the cost of accepting different culture decreases. First, minorities tend to accept different cultures, whereas the majority population tend to accept a different culture initially but reject it later. This is empirically supported by time series data on the number of foreign residents by nationality and municipality in Tokyo. Second, the number of firms producing minority-specific goods monotonically increases or shows an inverted U-shape. This is also empirically supported by cross-sectional data on the numbers of restaurants and residents by nationality and municipality in Tokyo.

Keywords: cultural acceptance, ethnic preference, diversity, assimilation, immigration
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1 Introduction

In the era of globalization, when migration across countries is common, societies are becoming more heterogeneous and diverse in ethnic composition. Indeed, “almost one in ten people living in OECD countries are foreign-born, and among younger cohorts (15- to 34-year-olds), over a quarter are foreign-born or native-born offspring of immigrant parents in OECD countries with available data, and the population shares of both groups have been increasing virtually everywhere (OECD, 2020, page 2).” Ethnic heterogeneity in a society, has various facets that cannot be found in an ethnically homogeneous society. For instance, ethnic heterogeneity in a society may induce friction among different ethnicities. As argued by Alesina and La Ferrara (2005), the costs of heterogeneity stem from the difficulty of reaching a decisive point or agreement on public good provisions and policies common to all ethnic groups. In such situations, mutual acceptance of diverse cultures among different ethnic groups would play a crucial role in mitigating conflicts among ethnic groups and stabilizing society.

In this paper, we consider the individual behavior of cultural acceptance in an ethnically diverse society. Since there are diverse levels of cultural acceptance, we take a revealed preference approach to two extreme levels of cultural acceptance. On the one hand, we focus on the consumption choice of ethnic-specific goods and services produced by other ethnic people as a light level of cultural acceptance. On the other hand, we consider the residential location choice by ethnic people as a deep level. In order to investigate the determinants of cultural acceptance, we build a model of ethnic preference where individuals decide how they consume other ethnic goods and housing located in a district with other ethnic people. In this paper, we interpret such consumption as a representation of cultural acceptance.

In our theoretical framework, there are two ethnic groups, majority and minority. Both ethnicities decide whether or not to consume other ethnic goods as well as their own ethnic goods.¹ We call this decision to consume other ethnic goods “accept other ethnicity (culture)” and decision not to consume other ethnic goods “reject other ethnicity (culture).” When accepting other ethnicity, an individual can consume more varieties of ethnic goods, which increases her utility. At the same time, she has to incur effort costs to get accustomed to other culture. The effort costs of accepting another ethnic culture are lower when the population size of the same ethnicity is larger. This is because information sharing or cooperation within the same ethnicity is considered helpful to individuals getting used to a society. We assume that the effort costs are inversely dependent on the ethnic population size. Our consideration of the effort costs is in line with the following literature. Epstein and Gang (2009) construct a model in which the minority’s efforts to assimilate with the

¹In the case of location choice, both ethnicities decide where to reside: segregated or integrated regions.
majority’s society and the degree to which the majority welcomes the minority play pivotal roles in helping minority people blend into the majority society. Lazear (1999) assumes that individuals in a smaller minority group value the importance of assimilation to the majority more than those in a larger minority group.

We show that both majority and minority individuals reject another ethnic culture when the effort costs are sufficiently high. By contrast, when the effort costs are low, either the majority or minority group accept the other ethnic culture. We also show that rising minority population induces minority individuals to accept the majority culture, while it induces majority individuals to accept first and then reject the minority culture. Similar things can be said for the changes in the number of ethnic firms based on our theoretical and numerical analysis for broad ranges of parameter values.

We conduct two empirical analyses by using municipality based data in the Tokyo prefecture. The first concerns the residential location patterns that are represented by segregation indices. This is because segregation indices are surrogates for the degree of assimilation/segregation, and thus, the degree of cultural acceptance. When examining the time series data on residential location of foreign residents by municipality, we find that foreign residents choose to assimilate to the majority society as the population of their own nationality increases. The second is on the spatial distribution of ethnic cuisine restaurants. Analyzing the cross-sectional data on the foreign restaurant distribution by municipality, we find a positive or inverted U-shaped relationship between the number of foreign residents and the number of ethnic cuisine restaurants. The two empirical analyses mostly support the theoretical results on assimilation to the indigenous society as well as the acceptance of different ethnic cultures.

The scope of our paper is related to the literature on ethnic diversity and assimilation/segregation. Examples of issues addressed in this literature include the following. Alesina and La Ferrara (2000) show that participation in social activities is significantly lower in less equal and more ethnically fragmented localities in the United States. Boustan (2007) points out the possibility that racial division reflects different preferences for public goods. She focuses on the civic costs of residing in an ethnically diverse jurisdiction, which is one of the factors explaining black-white segregation in the United States, where the costs come from compromising on the common public goods shared by different races. Chiswick et al. (2009) pay attention to ethnic-specific skills and examine the role

\footnote{Decreasing acceptance of the minority culture by majority individuals may be related to the US-China trade war and Brexit, which might be caused by increasing immigration.}

\footnote{According to OECD (2019), 475 thousand immigrants came to Japan in 2017, which is the fourth largest among OECD countries. The number of foreign residents in Japan has been gradually growing with an annual growth rate of 2.4% for 2006-2019.}
of complementarity or substitutability between them in determining the minority’s assimilation. Abdulloev et al. (2015) develop a model wherein consumption of a particular ethnic good hinders assimilation and provide empirical evidence supportive to it. Finally, Anas (2002) theoretically analyzes how racial prejudice and exclusion lead to residential segregation. Hársman (2006) shows that segregation in the Stockholm region has decreased among most of the 13 ethnic groups covered, but has increased significantly among Swedes, Iranians, and Africans. That is, rising ethnic diversity in the Stockholm region tends to increase minority individuals accepting Swedish culture, whereas there is a decrease in Swedes that accept minority culture. Previous literature has provided invaluable insights on assimilation and segregation, but has not fully considered individual decision on cultural acceptance, which is our main subject.

This paper is also related to the literature on cultural transmission (Bisin and Verdier, 2000). Studies in this literature mainly regard cultural transmission as a stochastic event. They also investigate factors that influence the probability of cultural transmission, and derive the resulting dynamics of the distribution of cultural traits in the population. Hence, cultural transmission occurs stochastically, which is regarded as exogenous for each individual. In contrast, we explicitly consider the effect of individual decision on cultural transmission (in our wording, acceptance of other ethnic culture). We believe that cultural transmission is partly determined stochastically and partly determined by individual decision. In this sense, this paper complements the existing literature on cultural transmission.

The paper is organized as follows. In the next section, we present a dynamic model of ethnic preferences and derive equilibrium and stability conditions. In order to get some insight, we obtain analytical solutions in a special case in Section 3 and conduct numerical analysis in Section 4. We then provide empirical evidences that support to our theoretical results in Section 5. Section 6 concludes.

## 2 Model of ethnic preference

We consider an economy which consists of two ethnic groups: the minority (foreign immigrants), indexed with $F$; and the majority (natives), indexed with $H$. The former population is $L_F$, while the latter is $L_H$, where $L_F < L_H$.  

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4See Bisin, Carvalho, and Verdier (2020) for a survey. Recent studies include Brueckner and Smirnov (2007), Patacchini and Zenou (2016), and Verdier and Zenou (2017).
2.1 Consumers

There are two types of differentiated goods associated to ethnicity $i \in \{F, H\}$: the minority good is called $F$-good and the majority good $H$-good. Examples include ethnic foods, dishes, clothing, arts, and music. Each individual with ethnicity $i$ consumes her own $i$-good. In addition, she decides whether or not to consume $j$-good with $j \neq i$ for $j \in \{F, H\}$.

If an individual chooses to consume a differentiated good of another ethnicity as well as her own ethnic good, we say that she accepts another ethnic culture ($e = A$). If she chooses not to consume another ethnic good but consumes her own ethnic good only, we say that she rejects another ethnic culture ($e = N$). We define $\lambda_i$ as a fraction of $i$-individuals who accept $j$-good ($e = A$), while $1 - \lambda_i$ is a fraction of $i$-individuals who reject $j$-good ($e = N$). We assume that individuals can freely move between the differentiated good sectors and provide their labor.

Individuals need to make some effort in order to understand another ethnic culture and enjoy consuming a good of the other ethnicity. The effort costs $k_i$ of ethnicity $i$ are assumed to be inversely proportional to its population size $L_i$. This reflects that sharing information for the adjustment and assimilation of different culture is easier for larger ethnic group. The effort costs are specified as

$$k_i = \frac{k \delta_i}{L_i},$$

where $k$ is a positive constant of effort costs and $\delta_i$ represents difficulties in accepting another ethnic culture. Individuals are heterogeneous in their difficulties; we assume that $\delta_i$ is uniformly distributed on the support of $[0, 1]$.

The utility function of an individual with ethnicity $i$ and cultural acceptance $e \in \{A, N\}$ is given by

$$U_{ie} = X_{ie} - I_e k_i,$$

$$X_{ie} \equiv \left[ \int_0^{n_i} x_i^j(v)^{\frac{\sigma - 1}{\sigma}} \, dv + I_e \int_0^{n_j} x_j^i(v)^{\frac{\sigma - 1}{\sigma}} \, dv \right]^{\frac{1}{\sigma - 1}}$$

for $j \neq i$, where

$$I_e = \begin{cases} 1 & \text{if } e = A \text{ (accept a good of another ethnicity)} \\ 0 & \text{if } e = N \text{ (reject a good of another ethnicity)} \end{cases}$$

is an indicator function, $n_i$ is the mass of firms producing $i$-good, $x_j^i(v)$ is the individual consumption of variety $v$ of $j$-good by ethnicity $i$, and $\sigma \, (> 1)$ is the elasticity of substitution between any two
varieties. The budget constraint of an ethnicity $i$-individual is

$$w = \int_{0}^{n_i} p_i(v)x^i_i(v)dv + I_e \int_{0}^{n_j} p_j(v)x^j_i(v)dv,$$

(3)

where $p_i(v)$ is the price of variety $v$ of $i$-good.

Maximizing the utility function (1) subject to the budget constraint (3) yields the following demand functions:

$$X_{ie} = \frac{w}{P_{ie}},$$

$$x^{jA}_i(v) = \frac{p_i^\sigma w}{P^{1-\sigma}_j}, \quad x^{jN}_i(v) = 0, \quad x^{iA}_i(v) = \frac{p_i^\sigma w}{P^{1-\sigma}_i}, \quad x^{iN}_i(v) = \frac{p_i^\sigma w}{P^{1-\sigma}_i},$$

where subscripts and superscripts $e \in \{A, N\}$ mean acceptance and rejection of a good of another ethnicity, respectively. The price index of ethnicity $i$ and cultural acceptance $e$ is defined by

$$P_{ie} \equiv \left[ \int_{0}^{n_i} p_i(v)^{1-\sigma}dv + I_e^\sigma \int_{0}^{n_j} p_j(v)^{1-\sigma}dv \right]^{\frac{1}{1-\sigma}}.$$  

(4)

### 2.2 Firms

Production of the differentiated good exhibits increasing returns to scale under monopolistic competition, so that each firm ends up producing a single variety. Production of each variety requires a marginal and a fixed labor requirements, $c$ and $f$, respectively. Each individual is endowed with one unit of labor and inelastically supplies it irrespective of nationality implying that labor is mobile between the $F$-good and $H$-good sectors. This leads to equalization of the wage rates between the sectors.\(^6\) We choose the labor as the numéraire, so that $w = 1$ holds true for all workers.

The profit of a firm producing variety $v$ in the $i$-good sector is given by

$$\pi_i(v) = p_i(v)q_i - w(cq_i + f),$$

\(^6\)Wage equalization does not hold in case firms producing ethnicity $i$-good employ workers with ethnicity $i$ only, which is in violation of the law. Nevertheless, it holds if like new economic geography there is a homogeneous good, which is produced by each ethnic group and traded freely.
where the demand for $i$-good of the firm is

$$q_i = \lambda_i L_i x_i^A(v) + (1 - \lambda_i) L_i x_i^N(v) + \lambda_j L_j x_j^A(v) + (1 - \lambda_j) L_j x_j^N(v)$$

$$= \lambda_i L_i \frac{p_i^\sigma w}{p_i^{1-\sigma}} + (1 - \lambda_i) L_i \frac{p_i^\sigma w}{p_i^{1-\sigma}} + \lambda_j L_j \frac{p_j^\sigma w}{p_j^{1-\sigma}}$$

for $i \neq j \in \{F, H\}$. \hfill (5)

Firm’s profit maximization yields

$$p_i = \frac{c \sigma w}{\sigma - 1},$$

$$P_{ie} = \frac{c \sigma w}{\sigma - 1} (n_i + I_e n_j)^{1-\sigma}. \hfill (6)$$

Plugging these prices into (5) with $w = 1$, we obtain the supply of $i$-good

$$q_i = \frac{\sigma - 1}{c \sigma} \left[ \lambda_i L_i (n_i + n_j)^{-1} + (1 - \lambda_i) L_i n_i^{-1} + \lambda_j L_j (n_i + n_j)^{-1} \right], \hfill (6)$$

and thus, the firm’s profit is

$$\pi_i(v) = \frac{1}{\sigma} \left[ \lambda_i L_i (n_i + n_j)^{-1} + (1 - \lambda_i) L_i n_i^{-1} + \lambda_j L_j (n_i + n_j)^{-1} \right] - f.$$

We assume free entry and exit of firms, which drives the firm’s profit to zero. Solving $\pi_F(v) = \pi_H(v) = 0$ yields the mass of firms in sector $i$ as follows:

$$n_i = \frac{L_i + L_j}{f \sigma} \frac{(1 - \lambda_i) L_i}{(1 - \lambda_i) L_i + (1 - \lambda_j) L_j}. \hfill (7)$$

From (7), we observe the following. First, $\partial n_i / \partial L_i > 0$ holds true simply because a larger population size $L_i$ yields a larger demand for $i$-good, resulting in a larger mass $n_i$ of $i$-good firms. Second, $\partial n_i / \partial \lambda_j > 0$ also holds. This is because as more $j$-individuals accept $i$-good, the demand for $i$-good increases, which raises the mass $n_i$ of $i$-good firms. Third, $\partial n_i / \partial (1 - \lambda_i) > 0$ holds because as more $i$-individuals reject $j$-good, the demand for $j$-good decreases, which results in a smaller mass $n_j$ of $j$-good firms. Since the total mass of firms is constant, i.e., $n_F + n_H = \frac{L_F + L_H}{f \sigma}$, a decline in $n_j$ implies that an increase in $n_i$. Thus, an increase in $(1 - \lambda_i)$ increases $n_i$.

Finally, the labor market clearing condition is satisfied in this economy because the aggregate labor supply $L_F + L_H$ is shown to be always equal to the aggregate labor demand $(c q_F + f) n_F + (c q_H + f) n_H$ with (6) and (7).
2.3 Equilibrium conditions

The indirect utility function is computed as

\[ V_{ie} = \frac{w}{P_{ie}} - I_e k_i, \]

which is expressed as

\[ V_{iA} = \frac{\sigma - 1}{c_\sigma} \left( \frac{L_i + L_j}{f_\sigma} \right)^{\frac{1}{\sigma - 1}} - \frac{k \lambda_i}{L_i}, \]  
\[ (8) \]

\[ V_{iN} = \frac{\sigma - 1}{c_\sigma} \left( \frac{L_i + L_j}{f_\sigma} \right)^{\frac{1}{\sigma - 1}} \left[ \frac{(1 - \lambda_i) L_i}{(1 - \lambda_i) L_i + (1 - \lambda_j) L_j} \right]^{\frac{1}{\sigma - 1}}. \]  
\[ (9) \]

Since the effort costs \( \delta_i \) are uniformly distributed over \([0, 1]\), individuals with \( \delta_i \) smaller (resp., larger) than the share \( \lambda_i \) of \( i \)-individuals accept (resp., reject) \( j \)-good. Therefore, we set \( \delta_i = \lambda_i \) in (8) and (9).

The equilibrium condition is given by \( V_{iA} = V_{iN} \), which is rewritten as

\[ \Delta V_i \equiv \frac{c_\sigma}{\sigma - 1} \left( \frac{f_\sigma}{L_i + L_j} \right)^{\frac{1}{\sigma - 1}} (V_{iA} - V_{iN}) = 0. \]  
\[ (10) \]

with a certain normalization. We use (10) in the rest of this paper. Several observations are in order. First, changes in \( \lambda_F \) or \( \lambda_H \) do not affect the total mass of varieties, which is given by \( n_F + n_H = \frac{L_F + L_H}{f_\sigma} \). Thus, changes in \( \lambda_F \) or \( \lambda_H \) do not alter the first term of \( V_{iA} \) in (8) since individuals consume all varieties when they accept another ethnicity. Second, changes in \( \lambda_F \) or \( \lambda_H \) affect the relative mass \( \frac{n_i}{n_i + n_j} \) of ethnic varieties. An increase in \( \frac{n_i}{n_i + n_j} \) improves \( V_{iN} \), which is reflected by the bracketed term in (9) such that \( \partial V_{iN}/\partial \lambda_i < 0 \) and \( \partial V_{iN}/\partial \lambda_j > 0 \). Suppose that \( \lambda_i \) increases, which intensifies the competition among firms producing \( i \)-good because \( i \)-individuals, who also consume \( j \)-good, would decrease consumption of \( i \)-good. Hence, it reduces the mass \( n_i \) of \( i \)-good firms as well as the utility \( V_{iN} \). On the other hand, it expands the market for varieties of ethnicity \( j \), which raises the mass \( n_j \) of \( j \)-good firms. Therefore, the increase in \( \lambda_i \) improves \( V_{jN} \).

Because of the nonlinear equilibrium conditions, there may exist multiple equilibria, some of which may be unstable. In order to refine such unstable equilibria, we consider the following dynamics of \( \lambda_i \) and examine the stability of the equilibria. Since an individual chooses to accept a good of another ethnicity \( (e = A) \) if \( \Delta V^i > 0 \) and does not \( (e = N) \) if \( \Delta V^i < 0 \), it would be
reasonable to employ the following simple dynamics:

\[ \frac{d\lambda_i}{dt} = \Delta V_i = 1 - \left[ \frac{(1 - \lambda_i)L_i}{(1 - \lambda_i)L_i + (1 - \lambda_j)L_j} \right]^{\frac{1}{\sigma-1}} - \frac{\sigma^{\frac{1}{\sigma-1}}K}{(\sigma - 1)(L_i + L_j)^{\frac{1}{\sigma-1}}} \lambda_i, \]  

for \( i \in \{F, H\} \). This is obtained by substituting (8) and (9) into (10). Even if an individual decision to accept a different culture takes place instantaneously, how it reflects on consumption will take some time. This may justify the introduction of the dynamics. Note that the composite cost parameter \( K \equiv cf^{\frac{1}{\sigma-1}}k \) involves marginal cost \( c \), fixed cost \( f \), and effort costs \( k \).

Since (11) is a two-variable dynamic, there are four corner solutions as candidates for equilibria. First, we exclude the full acceptance of both groups \( (\lambda_F, \lambda_H) = (1, 1) \). This is because the free-entry conditions collapse to \( n_F + n_H = \frac{L_F + L_H}{f} \), implying the indeterminacy of \( n_F \) and \( n_H \). Second, no acceptance of both groups \( (\lambda_F, \lambda_H) = (0, 0) \) is an equilibrium for infinite \( K \). As the sum of the first two terms of the RHS of (11) is positive and finite, the last term should be negative and finite when \( K \rightarrow \infty \). This is possible only when \( \lambda_i \rightarrow 0 \) for \( i \in \{F, H\} \). Finally, plugging \( (\lambda_F^*, \lambda_H^*) = (1, 0) \) and \( (0, 1) \) into the RHS of (11), and examining their signs, we can show that they are stable equilibria for small \( K \). In summation, we get the following results on corner equilibria.

**Proposition 1** Define \( K_F \equiv L_F(L_F + L_H)^{\frac{1}{\sigma-1}}(\sigma - 1)/\sigma^{\frac{1}{\sigma-1}} \) and \( K_H \equiv L_H(L_F + L_H)^{\frac{1}{\sigma-1}}(\sigma - 1)/\sigma^{\frac{1}{\sigma-1}} \).

If \( K \rightarrow \infty \), \( (\lambda_F^*, \lambda_H^*) = (0, 0) \) is a stable equilibrium.

If \( K_F < K \leq K_H \), \( (\lambda_F^*, \lambda_H^*) = (0, 1) \) is a stable equilibrium.

If \( K \leq K_F \), \( (\lambda_F^*, \lambda_H^*) = (0, 1) \) and \( (1, 0) \) are stable equilibria.

Proposition 1 states that the majority fully accept another good while the minority do not for intermediate \( K \), and that either the majority or minority fully accept another good for small \( K \). The latter is because low effort costs enable individuals to gain a high utility by consuming varieties provided by another ethnicity group.

Since thresholds \( K_F \) and \( K_H \) involve \( L_F \), Proposition 1 can be restated with respect to minority population size \( L_F \). It can be readily verified that when \( L_F = 0 \), \( (\lambda_F^*, \lambda_H^*) = (0, 0) \) is a unique stable equilibrium, leading to \( (n_F^*, n_H^*) = \left( 0, \frac{L_H}{f} \right) \). As \( L_F \) rises, both \( \lambda_i^* \) and \( n_i^* \) for \( i \in \{F, H\} \) initially increases. When \( L_F \) takes an intermediate value, which corresponds to \( K_F < K \leq K_H \), the effort costs of the minority group are so high that \( (\lambda_F^*, \lambda_H^*) = (0, 1) \) is a stable equilibrium, leading to \( (n_F^*, n_H^*) = \left( \frac{L_F + L_H}{f}, 0 \right) \) from (7). This implies that the demand for \( F \)-good is high because it is consumed by all individuals. On the other hand, demand for \( H \)-good is low because \( F \)-individuals do not consume \( H \)-good. As a result, \( H \)-good disappears from the economy. However, when \( L_F \) is sufficiently large, which corresponds to \( K \leq K_F \), the two corner solutions become stable equilibria.
with \((n^*_F, n^*_H) = \left( \frac{L_F + L_H}{f \sigma}, 0 \right), \left( 0, \frac{L_F + L_H}{f \sigma} \right)\). This suggests that when the population sizes of two ethnic groups do not differ much, multiple equilibria arise.

It should be noted that \((\lambda^*_F, \lambda^*_H) = (1, 0)\) is not a stable equilibrium for intermediate values of \(L_F\) in Proposition 1, whereas it is a stable equilibrium in the next three propositions. This is because Proposition 1 is only for corner equilibria. If we consider a stable equilibrium path taking interior equilibria into account, then \((\lambda^*_F, \lambda^*_H) = (1, 0)\) is a stable equilibrium while \((\lambda^*_F, \lambda^*_H) = (0, 1)\) never appears. Furthermore, such a stable equilibrium path is supported by empirical findings in Section 5.

Next, consider interior equilibrium solutions that can be obtained by solving the equilibrium conditions \(\Delta V_F = \Delta V_H = 0\), where we need to examine equilibrium stability of dynamics (11). In order to further refine equilibria, we consider a stable equilibrium path by a thought experiment of rising minority population \(L_F\) from zero or falling composite cost \(K\) from infinity.

Since dynamics (11) is highly nonlinear, it is difficult to fully characterize the stability of interior equilibria. Therefore, we focus on the special case of \(\sigma = 2\) in the next section and conduct a numerical analysis by changing various parameter values in Section 4.

### 3 Case of \(\sigma = 2\)

Suppose that the elasticity of substitution between two varieties is equal to two.

#### 3.1 Acceptance of different culture \((\lambda^*_F, \lambda^*_H)\)

We first consider the equilibrium share of people accepting different culture, \((\lambda^*_F, \lambda^*_H)\). We examine the stability of interior and corner solutions in order. Solving \(\Delta V_F = \Delta V_H = 0\) simultaneously in the range of \(0 < \lambda^*_i < 1\) for \(i \in \{F, H\}\), we obtain the following interior equilibrium solutions explicitly:

\[
(\lambda^*_F, \lambda^*_H) = \begin{cases} 
\left( \frac{L_F + L_H}{4K}, \frac{L_F + L_H}{4K} \right) & \text{for } K > \frac{L_F L_H}{4} \\
\left( \frac{L_F^2 + L_F L_H - 4K L_F}{4K(L_H - L_F)}, \frac{4K L_F^2 - L_F L_H}{4K(L_H - L_F)} \right) & \text{for } \frac{(L_F + L_H) L_F}{4} < K < \frac{(L_F + L_H) L_H}{4} \\
\end{cases}.
\] (12)

We can check their stability by computing eigenvalues of the Jacobians of dynamics (11) and evaluate them at equilibrium values. In addition, we know the stability of the corner equilibrium from Proposition 1. We provide a full stability analysis in Appendix A, which is summarized as follows.
Lemma 2 Suppose $\sigma = 2$.

(i) If $K > \frac{L_H(2F_H+L_H)}{4}$, there exists a unique stable equilibrium, $(\frac{L_FL_H}{4K}, \frac{L_HL_F}{4K})$.

(ii) If $\frac{L_HF}{2} < K \leq \frac{L_H(2F_H+L_H)}{4}$, there exist two stable equilibria, $(\frac{L_FL_H}{4K}, \frac{L_HL_F}{4K})$ and $(1,0)$.

(iii) If $K = \frac{L_HL_F}{2}$, there exists a unique stable equilibrium, $(1,0)$.

(iv) If $\frac{L_F(2F_H+L_H)}{4} < K < \frac{L_HL_F}{2}$, there exist two stable equilibria, $(\frac{(L_H^2+L_FL_H-4K)L_F}{4K(L_H-L_F)}, \frac{4K-L_F^2-L_FL_H}{4K(L_H-L_F)}L_H)$ and $(0,1)$.

(v) If $K \leq \frac{L_F(2F_H+L_H)}{4}$, there exist two stable equilibria, $(1,0)$ and $(0,1)$.

We know from Lemma 2 that $K$ and $L_F$ are key parameters that determine stable equilibrium. We seek a stable equilibrium path by thought experiments where we examine the effects of steady changes in these parameters on acceptance decisions of minority and majority people, $(\lambda^*_F, \lambda^*_H)$. The first thought experiment is falling $K$ from infinity to zero. As $K$ is a composite cost parameter, falling $K$ may be interpreted as technological progress over time. The second is rising minority population $L_F$. Such increasing immigration has been observed in recent years due to globalization, as mentioned in the introduction. It should be noted that setting the initial condition ($K$ infinity in the former and $L_F$ zero in the latter) is another refinement of equilibrium to make the equilibrium unique. We readily have the following propositions.

Proposition 3 Given $\sigma = 2$, the stable equilibrium path for falling composite cost $K$ is as follows.

(i) From $+\infty$ to $\frac{L_HL_F}{2}$, both $\lambda^*_F$ and $\lambda^*_H$ increase.

(ii) From $\frac{L_HL_F}{2}$ to $\frac{L_F(2F_H+L_H)}{4}$, $\lambda^*_F$ increases to 1 while $\lambda^*_H$ decreases to 0.

(iii) From $\frac{L_F(2F_H+L_H)}{4}$ to 0, corner equilibrium $(\lambda^*_F, \lambda^*_H) = (1,0)$ continues.

We know from this proposition that $\lambda^*_F$ always rises, while $\lambda^*_H$ shows an inverted U-shape with respect to $K$. That is, technological progress caused by falling $K$ always makes the minority group accept a majority good, whereas it makes the majority group accept a minority good only in the first half of the inverted U-shaped curve.

Proposition 4 Given $\sigma = 2$, the stable equilibrium path for rising minority population $L_F$ is as follows.

(i) From 0 to $2K/L_H$, both $\lambda^*_F$ and $\lambda^*_H$ increase.

(ii) From $2K/L_H$ to $\frac{\sqrt{L_H^2+16K-L_H}}{2}$, $\lambda^*_F$ increases to 1 while $\lambda^*_H$ decreases to 0.

(iii) From $\frac{\sqrt{L_H^2+16K-L_H}}{2}$ to 1, corner equilibrium $(\lambda^*_F, \lambda^*_H) = (1,0)$ continues.

---

1Because we use asymptotic stability, we cannot determine the stability of the interior equilibrium when $K = \frac{L_HL_F}{2}$. If we employ Lyapunov stability, the interior equilibria are stable when $K = \frac{L_HL_F}{2}$, implying that there exist two stable equilibria.
Again, $\lambda^*_F$ always rises, while $\lambda^*_H$ is an inverted U-shape with respect to $L_F$. Economic interpretations are stated at the end of the next section.

The blue dotted piecewise linear line in Figure 1 plots the stable equilibrium path $(\lambda^*_F, \lambda^*_H)$ when $L_F$ steadily increases from zero to $L_H$ for $\sigma = 2$, $L_H = 3$, and $f = 1$,\textsuperscript{8} which implies $K = ck$: that is, changing the value of $K$ is equivalent to changing the value of $k$ or $c$.

![Figure 1: The effect of increasing $L_F$ on the acceptance of different culture $(\lambda^*_F, \lambda^*_H)$, where the blue dotted curve is for $\sigma = 2$, green solid curve is for $\sigma = 4$, and red dashed curve is for $\sigma = 9$.](image)

We can draw the same path for falling $K$ from infinity to zero. To summarize, the impacts of lowering the composite cost on $\lambda^*_F$ and $\lambda^*_H$ are qualitatively the same as those of increasing the minority population.

### 3.2 The mass of firms $(n^*_F, n^*_H)$

Next, we examine changes in the mass of firms providing ethnic goods, $(n^*_F, n^*_H)$. Substituting the stable equilibrium $(\lambda^*_F, \lambda^*_H)$ into (7) yields the stable equilibrium mass of firms. The interior stable equilibrium of firms is given by

$$\left(n^*_F, n^*_H\right) = \begin{cases} \left(\frac{L_F}{2f}, \frac{L_H}{2f}\right) & \text{for } K > \frac{L_F L_H}{2} \\ \left(\frac{L_H^2 - L_F L_H - 4K}{2f(L_H - L_F)}, \frac{L_H^2 + L_F L_H - 4K}{2f(L_H - L_F)}\right) & \text{for } \frac{(L_F + L_H)L_F}{4} < K < \frac{L_F L_H}{2}. \end{cases} \tag{13}$$

The stable corner equilibrium $(\lambda^*_F, \lambda^*_H) = (1, 0)$ is translated into $(n^*_F, n^*_H) = \left(0, \frac{L_F + L_H}{f}\right)$. Therefore, Proposition 4 can be rewritten with respect to $(n^*_F, n^*_H)$ as follows.

**Proposition 5** Given $\sigma = 2$, the stable equilibrium path for rising minority population $L_F$ is as follows.

(i) From 0 to $2K/L_H$, $n^*_F$ increases while $n^*_H$ does not change.

\textsuperscript{8}The parameter values of $L_H = 3$ and $f = 1$ are common to all the panels of Figures 1 and 2.
(ii) From $\frac{2K}{L_H}$ to $\sqrt{\frac{L_H^2 + 16K - L_H}{2}}$, $n^*_F$ decreases to 0 while $n^*_H$ increases.

(iii) From $\sqrt{\frac{L_H^2 + 16K - L_H}{2}}$ to 1, $n^*_F = 0$ while $n^*_H$ increases.

Hence, we confirm that the mass $n^*_H$ of firms producing majority varieties is weakly increasing with respect to $L_F$ as $\lambda^*_F$ is. On the other hand, the mass $n^*_F$ of firms producing ethnic varieties forms an inverted U-shape with respect to $L_F$ as $\lambda^*_H$ does. Again, their economic interpretations are stated at the end of this section.

![Graphs](image)

Figure 2: The effect of increasing $L_F$ on the number of $n^*_F$ of restaurants with different sets of $(\sigma, K, L_H)$, where blue dotted curves are for $\sigma = 2$, green solid curves are for $\sigma = 4$, and red dashed curves are for $\sigma = 9$.

The blue dotted curve in each panel of Figure 2 displays the impact of $L_F$ on $n^*_F$ under different parameter sets of $(L_H, K)$ when $\sigma = 2$. It verifies Proposition 5: the mass $n^*_F$ of firms producing ethnic varieties has an inverted U-shaped or monotonically increasing relationship with minority population $L_F$. Comparison of blue dotted curves among the panels in Figure 2 also shows that the mass $n^*_F$ of firms producing ethnic varieties is larger for larger composite cost $K$.

To summarize, we can say that, from Propositions 4 and 5, $\lambda^*_H$ and $n^*_F$ are inverted U-shapes with respect to $L_F$, whereas $\lambda^*_F$ and $n^*_H$ are monotone increasing in $L_F$. These results may be intuitively understood as follows. As minority population $L_F$ steadily increases, effort costs $k_F$ decrease because sharing information for adjustment and assimilation of different cultures is easier.
for a larger ethnic population. Hence, $\lambda^*_F$ increases, which raises consumption of $H$-good, and thus, $n^*_H$ also increases.

In contrast, the inverted U-shaped path of $\lambda^*_H$ with respect to $L_F$ depends on positive and negative effects. The positive effect of $L_F$ on $\lambda_H$ is the following. Rising $L_F$ increases $n_F$ as we showed at the end of Section 2.2. The increase in $n_F$ encourages majority individuals to accept the minority culture, i.e., $\lambda^*_H$ increases. On the other hand, the negative effect of $L_F$ on $\lambda_H$ is as follows. Increasing $L_F$ raises $\lambda^*_F$ due to reductions in the effort costs for minority individuals. This decreases $1 - \lambda_F^*$, which reduces $n_F$ as we showed at the end of Section 2.2. Then, majority individuals would reject the minority culture in order to avoid the effort costs, and hence, $\lambda^*_H$ decreases. The positive effect would be dominant in the first phase of increasing $\lambda^*_H$, whereas the negative effect would be stronger in the second phase of decreasing $\lambda^*_H$.

We will check the robustness of these results against alternative values of $\sigma$ in the next section.

4 Numerical analysis

Figure 1 also depicts the impact of minority population $L_F$ on $(\lambda^*_F, \lambda^*_H)$ by numerical simulations for $\sigma = 4$ and 9. Acceptance share $\lambda^*_F$ of different culture by minority is always increasing in $L_F$, whereas acceptance share $\lambda^*_H$ of a different culture by the majority forms an inverted U-shape followed by a sudden jump to the horizontal axis $\lambda^*_H = 0$, which continues for large $L_F$.

Analogously, Figure 2 plots the impact of $L_F$ on $n^*_F$ by numerical simulations. To summarize, the mass $n^*_F$ of firms providing ethnic varieties exhibits either (A) a monotonically increasing curve with respect to minority population $L_F$ or (B) an inverted U-shaped curve and then a sudden jump to the horizontal axis $n^*_F = 0$, which continues for large $L_F$. It is noted that the U-shape and the jump path in (B) is observed not only for $n^*_F$ but also for $\lambda^*_H$.

These findings are verified by extensive numerical simulations for all combinations of nine values of $\sigma$ from 2 to 10, six values of $K$ from 0.1 to 2, and five values of $K$ from 2 to 15. Thus, the overall results that $n^*_F$ is either monotonically increasing or an inverted U-shape in $L_F$ are considerably robust. These theoretical results are empirically tested in Section 5 by using municipality based data on the population share by nationality and the number of ethnic restaurants.

\[^9\text{Numerical results are available upon request.}\]
\[^{10}\text{When } \sigma < 2, n^*_F \text{ of the vertical axis is replaced with } n^*_H \text{ according to numerical simulations. However, we focus only on the case of } \sigma > 2 \text{ following the estimates of } \sigma \text{ in the literature.}\]
5 Empirical findings

In this section, we provide some empirical evidence that supports the theoretical results obtained in the previous sections. First, we look at the assimilation/segregation patterns represented by residential choice by foreign and Japanese residents as a deep level of cultural acceptance in Section 5.1. This is a time series analysis using municipality based data of foreign residents in Tokyo for 1954-2019. Throughout the paper, we exclude towns and villages in Tokyo because there are few foreign residents in these regions. The degree of assimilation of individuals of ethnic minorities to the majority’s society depends on their identity choices.\textsuperscript{11} If they strongly identify with their own ethnic group, they would reject the majority’s social norms, resulting in ethnic segregation. If their identification with their own ethnic group is not strong, they choose to, at least in part, assimilate to the majority’s norm, resulting in a higher level of assimilation. Moreover, as shown by Atkin et al. (2020), individual identity choices are reflected by consumption behavior.\textsuperscript{12} Hence, we interpret the share $\lambda^*_i$ of individuals accepting a good of another ethnicity as the degree of assimilation, so that we empirically investigate our theoretical results by looking at assimilation/segregation patterns, which result from consumption choice of housing, i.e., locational choice of housing by different ethnic groups.

Second, we pay attention to the acceptance of foreign cultures by looking at the spatial distributions of foreign restaurants and residents as a light level of cultural acceptance in Section 5.2. We conduct a cross sectional analysis using municipality based data of foreign restaurants and residents in Tokyo for 2020, motivated by Mazzolari and Neumark (2012), who examined the effects of immigration on the diversity of consumption choices with restaurant data in California. They showed that immigration is associated with increased ethnic diversity of restaurants. Hence, it would be reasonable to focus on restaurants as a proxy for stores providing ethnic goods. Our model predicts that the number of firms providing ethnic varieties, $n^*_i$, is determined by $\lambda^*_i$ and hence, by $L_i$, where the relationship between $n^*_i$ and $L_i$ is shown in Figure 2. We examine whether this relationship is observed in our restaurant data.

5.1 Assimilation and segregation

First, we study the assimilation/segregation transition in Tokyo prefecture. As explained in Appendix A, the data on the number of foreign residents are missing in some years for 10 nationalities.\textsuperscript{11} This view is shared by existing studies on assimilation such as Battu et al (2007), who investigated the labor market assimilation from the perspective of ethnic minority’s identity choice.\textsuperscript{12} Atkin et al. (2020) focused on India, which is characterized by deep ethnic and linguistic divisions, finding that consumption of specific goods (e.g., beef and pork) responds to the degree of identification with a particular social group.
Furthermore, the number of municipalities has been changing between 30 and 50 due to municipal merger and division during the study period. Keeping these in mind, we calculate the dissimilarity indices below and check whether they fit the theoretical findings obtained in the previous sections. During the study period of 1954-2019 in Tokyo wards and cities, the annual growth rate of foreign residents was 3.32, while that of Japanese residents was 0.95. Therefore, the time series changes in these indices would be mainly affected by increasing foreigners rather than Japanese.

The index of dissimilarity is by Duncan and Duncan (1955). It describes the overall extent of uneven distribution of nationality $i$-residents and nationality $H$-residents,

\[
ID_i \equiv \frac{1}{2} \sum_{r=1}^{m} \left| \frac{L_{ir}}{L_i} - \frac{L_{iH}}{L_H} \right| \quad \text{for } i \neq H,
\]

where $L_{ir}$ is the total population of nationality $i$ in municipality $r$, $L_i$ is the total population of nationality $i$ in Tokyo, $m$ is the number of municipalities, $i =$ Chinese, Korean, Vietnamese, Filipino, Nepali, American, Indian, Burmese, and Thai, while $H =$ Japanese. I should be noted that since the dissimilarity is not by race, but by nationality. Index $ID_i$ coincides with the Hoover index if the last term is replaced with the area share in municipality $r$. This index is small if the spatial distribution of nationality $i$-residents is similar to that of Japanese residents, and thus, represents how nationality $i$-residents assimilate and integrate with Japanese residents.

Let $L_r$ be the total population in municipality $r$ and $L \equiv \sum_{r=1}^{m} L_r$ be the total population in Tokyo. The second index is

\[
DI_i \equiv \frac{L}{2L_i (L - L_i)} \sum_{r=1}^{m} L_r \left| \frac{L_{ir}}{L_r} - \frac{L_i}{L} \right|,
\]

which is called the dissimilarity index (Massey and Denton, 1988). The third index $CV_i$ is the coefficient of variation of the share of nationality $i$-residents in municipality $r$, $L_{ir}/L_r$. These three indices are qualitatively similar in that the smaller (resp., larger) the values of the indices are, the more assimilated (resp., segregated) nationality $F$-residents are.

After calculating these indices, we regress them on year $t$ in order to see the time trend. Then, we examine the sign and $t$-value of the regression coefficient of the year. If it is significantly positive (resp., negative), then country $i$-residents are getting more dissimilar (resp., similar) and resides more unevenly (resp., evenly) in Tokyo prefecture. The results are tabulated in Table 1. The table shows that the signs are the same and their significance is almost the same between $ID_i$ and $DI_i$, and

\[13\]Our classification of foreign nationalities follows the Statistics on Foreign Residents, the Ministry of Justice, Japan, in which the foreign nationalities are labeled as “foreign countries or regions.” For example, both immigrants and their descendants from the Korean Peninsula have been categorized under the Korean nationality.
Table 1: t-values of the time-trend slopes of the dissimilarity indices

<table>
<thead>
<tr>
<th>Nationality</th>
<th>t-value of slope of ID</th>
<th>t-value of slope of DI</th>
<th>t-value of slope of CV</th>
</tr>
</thead>
<tbody>
<tr>
<td>Japan</td>
<td>-</td>
<td>+7.89***</td>
<td>+4.39***</td>
</tr>
<tr>
<td>China</td>
<td>-3.01***</td>
<td>-3.34***</td>
<td>-0.56</td>
</tr>
<tr>
<td>Korea</td>
<td>+7.76***</td>
<td>+6.72***</td>
<td>+9.97***</td>
</tr>
<tr>
<td>Philippines</td>
<td>-7.64***</td>
<td>-7.74***</td>
<td>-6.46***</td>
</tr>
<tr>
<td>USA</td>
<td>-6.36***</td>
<td>-6.50***</td>
<td>+1.49*</td>
</tr>
<tr>
<td>India</td>
<td>+1.60*</td>
<td>+1.55*</td>
<td>+1.88**</td>
</tr>
<tr>
<td>Thailand</td>
<td>-7.33***</td>
<td>-7.63***</td>
<td>-6.88***</td>
</tr>
</tbody>
</table>

* = significant at 10% level; ** = significant at 5% level; *** = significant at 1% level

that the signs and their significance of CV$_i$ are similar to that of ID$_i$ and DI$_i$. The data indicate that Japanese, Korean, and Indian residents have been segregating, whereas Chinese, Filipino, American, and Thai residents have been assimilating during the study period of 1954-2019. Time series plots of DI$_i$ for Japanese, Chinese and Korean residents are shown in Figure 3. Plots of DI$_i$ for other nationalities, and ID$_i$ and CV$_i$ for all nationalities are contained in Figure A1-A3 of Appendix C.

We can relate these empirical results to analytical ones in Section 3 as follows. First, segregating Japanese residents corresponds to decreasing $\lambda_H^*$ in Proposition 4(ii), which is the right-half phase of the inverted U-shaped curve. Second, assimilating residents of Chinese, Filipino, American, and Thai residents corresponds to increasing $\lambda_F^*$ in Proposition 4(i) and (ii), which is somewhere on the inverted U-shaped curve. Putting these two observations together, we may conclude that immigration from China, Philippines, USA, and Thailand are on the right half phase of the inverted U-shaped curve.

Next, since the data in this section are panel data, we test whether these indices are smaller for larger immigrants by nationality. Let $L_F(t)$ be the number of residents of nationality $F$ in year $t$,
and $ID_F(t)$, $DI_F(t)$ and $CV_F(t)$ are the three indices of nationality $F$ in year $t$. Then, combining the nationality and year, there are 237 observations. The correlation of $L_F(t)$ with $ID_F(t)$ is $-0.513$ and its $t$-value is $-9.17$, that with $DI_F(t)$ is $-0.513$ and its $t$-value is $-9.16$, and that with $CV_F(t)$ is $-0.495$ and its $t$-value is $-8.72$. Thus, we can conclude that immigrants tend to be assimilated as they increase.

It should be fair to mention the segregation tendency of Korean and Indian residents, which corresponds to decreasing $\lambda^*_F$. Such a decrease does not appear in Proposition 4, suggesting that there are some other reasons that are not incorporated in our model.

First, the segregation tendency of Korean residents may be mainly attributed to the fact that 5,000 to 10,000 Korean people residing in Japan become naturalized as Japanese every year.\(^{14}\) They are likely to assimilate to the society due to small effort costs, whereas non-naturalized Korean residents would incur large effort costs. Since the statistics of Korean nationalities that we are using involve the latter only, it is no surprise that Korean residents exhibit the tendency to segregate.

Second, the tendency to segregate of Indian residents may be due to the fact that Indian residents in Japan are more likely to live with their families than other nationality immigrants do. According to Sawa and Minamino (2009), “Dependent Visa” ranks top in the status of residence for Indians unlike other foreign residents in Japan. It may be that Indian residents are not eager to communicate every day with native Japanese residents in order to maintain their mental well-being, as they feel less isolated when living with their families. Since the theoretical model does not capture such mental benefits from communication into the utility function, the segregation tendency of Indian residents may not fit the results of our model.

5.2 Immigration and spatial distribution of restaurants

We next examine the spatial distribution of restaurants serving foreign foods and dishes. We obtained data on the number of residents by nationality and municipality in Tokyo prefecture for 2020 from the Tokyo Statistical Yearbook (https://www.toukei.metro.tokyo.lg.jp/gaikoku/2020/ga20010000.htm). We consider nine foreign nationalities having large populations; $F = \text{Chinese, Korean, Vietnamese, Filipino, Nepali, American, Indian, Burmese, and Thai}$, while $H = \text{Japanese}$ as in the previous section. Table 2 lists the number and share of foreign residents by nationality in Tokyo prefecture.

The share of foreign residents in Tokyo prefecture is 4.09%, which is larger than that in the rest of Japan. Some municipalities have much larger shares: 12% in Shinjuku Ward and 9.5% in

\(^{14}\) More precisely, 55.6% of the total change in Korean population in Japan is attributed to the population change by naturalization on average during the period between 1985 and 2013. This exceeds the immigration from Korea (29.6%) and the natural change (14.8%) according to MINDAN database (https://www.mindan.org/syakai.php).
Table 2: Foreign residents of major nationalities on January 1, 2020 in Tokyo

<table>
<thead>
<tr>
<th>Nationality</th>
<th># of residents</th>
<th>% share in total population</th>
<th>% share in foreign residents</th>
</tr>
</thead>
<tbody>
<tr>
<td>Total (Japanese + Foreign)</td>
<td>13,942,856</td>
<td>100</td>
<td>-</td>
</tr>
<tr>
<td>Total (Foreign)</td>
<td>570,165</td>
<td>4.09</td>
<td>100</td>
</tr>
<tr>
<td>Chinese</td>
<td>226,976</td>
<td>1.63</td>
<td>39.8</td>
</tr>
<tr>
<td>Korean</td>
<td>92,241</td>
<td>0.66</td>
<td>16.2</td>
</tr>
<tr>
<td>Vietnamese</td>
<td>37,377</td>
<td>0.27</td>
<td>6.6</td>
</tr>
<tr>
<td>Filipino</td>
<td>34,071</td>
<td>0.24</td>
<td>6.0</td>
</tr>
<tr>
<td>Nepali</td>
<td>25,832</td>
<td>0.19</td>
<td>4.5</td>
</tr>
<tr>
<td>American</td>
<td>19,154</td>
<td>0.14</td>
<td>3.4</td>
</tr>
<tr>
<td>Indian</td>
<td>13,998</td>
<td>0.10</td>
<td>2.5</td>
</tr>
<tr>
<td>Burmese</td>
<td>10,008</td>
<td>0.07</td>
<td>1.8</td>
</tr>
<tr>
<td>Thai</td>
<td>7,969</td>
<td>0.06</td>
<td>1.4</td>
</tr>
</tbody>
</table>

Total population in Tokyo Prefecture on October 1, 2019. Foreign population by nationality on January 1, 2020 in Tokyo.

Toshima Ward. In these municipalities, the consumption behavior of foreign residents can have significant impacts on the spatial distribution of restaurants, which in turn affects the consumption behavior of native people.

*Tabelog* (https://tabelog.com/) is a website that provides restaurant information in Japan. Like Yelp (https://www.yelp.com/), it is a review site, but it specializes in restaurant information. It is one of the most popular review sites: the number of registered restaurants exceeds 0.9 million, and the number of reviews exceeds 3.5 millions (https://tabelog.com/help/beginner/ as of May 12, 2020). Users with *Tabelog* accounts and owners of restaurants can register restaurants on the site. They can also make comments and rate restaurants. Regardless of the registration status, users can search restaurants and see the registered information (menu, comments, rates, etc.). We collected data on the number of restaurants serving dishes by nationality listed in Table 2 and by municipality in Tokyo prefecture. See Appendix B for explanations of the data.

The restaurant list of *Tabelog* is based on registration by users and owners rather than an exhaustive survey. Nevertheless, given its most popularity in Japan, it certainly has an almost comprehensive list of restaurants in Japan. Thus, the number of restaurants serving cuisines by foreign nationality reflects how the ethnic cuisine culture would be accepted in Japan.

Figure 4 plots the data for Chinese and Korean cuisines, whose population sizes in Tokyo prefecture are the largest. The horizontal axis $L_{Fr}/L_{Hr}$ is the number of nationality $F$-residents per one thousand Japanese people in municipality $r$ and the vertical axis $n_{Fr}/L_{Hr}$ is the number of...
Figure 4: The numbers of foreign residents and foreign restaurants by municipality, where the horizontal axis is \( \frac{10^3 L_{Fr}}{L_{Hr}} \) and the vertical axis is \( \frac{10^3 n_{Fr}}{L_{Hr}} \).

nationality (cuisine) \( F \)-restaurants serving dishes per one thousand Japanese people in municipality \( r \). As before, we exclude towns and villages, where there are few or no restaurants. The number of municipalities in 2020 is \( m = 49 \). The curve is fitted by a quadratic function passing through the origin in each panel of Figure 4.

The fitted curve is an inverted U-shape in panel 4a for Chinese cuisine, whereas it is monotonically increasing in panel 4b for Korean one. Plots of the seven other cuisines appear in Figure A4 in Appendix C. We observe that the cross sectional variations look large, possibly due to the assumption of autarky within each municipality. That is, residents do not cross over municipality borders. In reality, however, residents often visit and work in restaurants outside their municipality. In fact, the average pairwise distance between two municipal city halls is 20km ranging from 1.6km and 54km, suggesting that residents would often visit and work in restaurants outside their municipalities. Therefore, we take inter-municipality movements into consideration below.

First, we employ the market potential à la Harris (1954) instead of the indices of two axes in Figure 4. This is because such visit and work would exhibit a distance-decay property: people visit more distant places less frequently. That is, we replace \( L_{Fr}/L_{Hr} \) and \( n_{Fr}/L_{Hr} \) with the resident and restaurant potentials, respectively, defined as

\[
\frac{PL_{Fr}}{PL_{Hr}} \quad \text{and} \quad \frac{Pn_{Fr}}{PL_{Hr}},
\]

where

\[
PL_{ir} = \sum_{s=1}^{m} \frac{L_{is}}{\exp (\tau d_{rs})} \quad \text{and} \quad Pn_{ir} = \sum_{s=1}^{m} \frac{n_{is}}{\exp (\tau d_{rs})} \quad \text{for} \ i \in \{ F, H \},
\]

\( d_{rs} \) is the geographic distance between municipal city halls \( r \) and \( s \), and \( \tau \) is a parameter of distance friction. Second, intra-municipal distance is given by \( d_{rr} = \sqrt{S_r/\pi}/2 \), where \( S_r \) is the total
area in municipality \( r \). Third, we take commuting into account. Many people living in Tokyo prefecture work in the central city of Tokyo, which consists of three municipalities: Chiyoda, Chuo, and Minato Wards, indexed \( r = 1, 2, 3 \), respectively. People go to restaurants in the central city during weekdays and to restaurants in their own municipality during weekend. Assuming that the consumer’s probability of the former is \( \rho \) and that of the latter \( 1 - \rho \), and that the worker’s probabilities of finding a job in the central city and in their own municipality are also the same, then the potential (14) can be rewritten as

\[
\rho \sum_{s=1}^{3} \frac{P_{L_{Fs}}}{PL_{Hs}} + (1 - \rho) \frac{P_{L_{Fr}}}{PL_{Hr}} \text{ and } \rho \sum_{s=1}^{3} \frac{P_{n_{Fs}}}{PL_{Hs}} + (1 - \rho) \frac{P_{n_{Fr}}}{PL_{Hr}}
\]

(15)

for \( r = 4, 5, \ldots, 49 \), while (14) remain unchanged for the central municipalities \( r = 1, 2, 3 \). Finally, we take ad hoc values of the parameters \( \tau = 1 \) and \( \rho = 1/5 \).16

With these adjustments of indices, we revised Figure 4 to Figure 5.

Figure 5: The adjusted numbers of foreign residents and foreign restaurants by municipality, where the horizontal axis is \( \rho \sum_{s=1}^{3} \frac{P_{L_{Fs}}}{PL_{Hs}} + (1 - \rho) \frac{P_{L_{Fr}}}{PL_{Hr}} \) and the vertical axis is \( \rho \sum_{s=1}^{3} \frac{P_{n_{Fs}}}{PL_{Hs}} + (1 - \rho) \frac{P_{n_{Fr}}}{PL_{Hr}} \).

Observe that the quadratic function in Figure 5 fits better, exhibiting an inverted U-shape in Chinese cuisine and a monotonically increasing one in the Korean case. Figure A5 in Appendix C contains plots for the seven other countries, some of which are an inverted U-shape, such as Chinese, and others are increasing, such as Korean.

In order to represent this statistically, we conduct the \( t \)-test developed by Lind and Mehlum (2010), which tests the presence of an inverted U-shaped relationship between the two indices in (15). If the slope of the quadratic function at the minimum value of the first index in (15) is significantly 

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15 This formula is due to the assumption that the production in each region is concentrated in a single point at the center of the disk, where the density of consumers linearly decreases in relation to the distance from the center.

16 This is simply because the quadratic function fit well by these parameter values, although changing them does not cause them to differ much.
positive and that at the maximum value of the first index in (15) is significantly negative, then the relationship is called an inverted U-shape. If both of the slopes are significantly positive, then the relationship is called increasing. The results are listed in Table 3.

Table 3: t-values of the slopes at the minimum and maximum values

<table>
<thead>
<tr>
<th>Nationality</th>
<th>$R^2$</th>
<th>Estimated equation by OLS</th>
<th>t-value at minimum</th>
<th>t-value at maximum</th>
</tr>
</thead>
<tbody>
<tr>
<td>Chinese</td>
<td>0.895</td>
<td>$0.125x - 0.00236x^2$</td>
<td>+19.8***</td>
<td>-2.69***</td>
</tr>
<tr>
<td>Korean</td>
<td>0.911</td>
<td>$0.0382x - 0.000004x^2$</td>
<td>+21.3***</td>
<td>+3.72***</td>
</tr>
<tr>
<td>Vietnamese</td>
<td>0.775</td>
<td>$0.0216x - 0.00181x^2$</td>
<td>+10.5***</td>
<td>-1.56*</td>
</tr>
<tr>
<td>Filipino</td>
<td>0.645</td>
<td>$0.00181x - 0.000281x^2$</td>
<td>+8.85***</td>
<td>-1.27</td>
</tr>
<tr>
<td>Nepali</td>
<td>0.836</td>
<td>$0.0421x - 0.00450x^2$</td>
<td>+12.7***</td>
<td>-2.53***</td>
</tr>
<tr>
<td>American</td>
<td>0.973</td>
<td>$0.0239x - 0.000123x^2$</td>
<td>+35.9***</td>
<td>+7.33***</td>
</tr>
<tr>
<td>Indian</td>
<td>0.940</td>
<td>$0.221x - 0.0334x^2$</td>
<td>+26.5***</td>
<td>-9.96***</td>
</tr>
<tr>
<td>Burmese</td>
<td>0.882</td>
<td>$0.00273x + 0.000501x^2$</td>
<td>+6.18***</td>
<td>+7.92***</td>
</tr>
<tr>
<td>Thai</td>
<td>0.867</td>
<td>$0.278x - 0.0898x^2$</td>
<td>+17.4***</td>
<td>-0.28</td>
</tr>
</tbody>
</table>

* = significant at 10% level; ** = significant at 5% level; *** = significant at 1% level

The t-value at each minimum value is significantly positive at the 1% level, whereas that at each maximum value is positive or negative. It is significantly negative for Chinese, Nepali, and Indian cuisines at the 1% level and Vietnamese cuisine at the 10% level, while it is insignificantly negative in the Filipino and Thai cases. These cuisines exhibit an inverted U-shaped curve: the relationship is positively related for few immigrants, but negatively related for many immigrants. Such a pattern is consistent with the relationship between immigrant population $L_F$ and the number $n^*_F$ of restaurants shown in Proposition 5 as well as the simulation results of (B) inverted U-shaped curves in Figure 2.

On the other hand, the t-value at each maximum value is significantly positive for Korean, American, and Burmese cuisines, which means an increasing curve: the number of restaurants is positively related to the immigrant population. This is also consistent with the simulation results of (A) monotonically increasing in panels (b) and (d) in Figure 2.

6 Conclusion

We have developed a model of ethnic preference for different culture and analyzed the cultural behaviors of ethnic minority and majority people. First, we have shown in Proposition 4(ii) that as the minority people steadily increases (or the composite cost steadily decreases), acceptance of different culture keeps increasing in the case of minority people, whereas initially rises and then falls (i.e., inverted U-shaped) in the case of majority people. Second, we have shown in Proposition
that as the minority people increases (or the composite cost steadily decreases), the number of firms producing ethnic varieties monotonically increases or exhibits an inverted U-shaped curve.

We then conducted an empirical analysis in order to test these theoretical results. First, using the time series data on the number of foreign residents by nationality and municipality, we have shown the segregating tendency of Japanese residents and assimilating tendency of Chinese, Filipino, American, and Thai residents, which support Proposition 4(ii) and the simulation results, i.e., the right half of the inverted U-shaped curve. Second, using the cross sectional data on the number of restaurants and the number of foreign residents by nationality and municipality, we observed a monotonically increasing relationship between these two numbers for Korean, American, and Burmese cuisines, while an inverted U-shaped relationship for Chinese, Nepali, Indian, and Vietnamese ones. This would justify Proposition 5 as well as the simulation results.

Our results are significant in designing assimilation policies. Both minority and majority people are likely to accept a different culture in the early stages of consecutive increases in immigration, whereas majority people come to reject different culture when the immigration size becomes large. In such a circumstance, education policies fostering the understanding of the majority people regarding minority culture would be called for.
References


Appendix A: Stability of equilibrium

We first examine the stability of interior equilibria. Solving $\Delta V_F = \Delta V_H = 0$ simultaneously in the range of $0 < \lambda^*_i < 1$ for $i \in \{F, H\}$, we obtain the interior equilibrium solutions given by (12). In order to check their stability, we compute the eigenvalues of the Jacobians of dynamics (11) and evaluate them at the equilibrium values. The eigenvalues evaluated at the first equilibrium of (12) are

$$2K \left[ L_H L_F \left( L_F^2 + 4L_F L_H + L_H^2 \right) - 4K (L_F + L_H)^2 \pm \sqrt{(4K - L_F L_H)^2 (L_H^2 - L_F^2)^2 + 4L_F^4 L_H^4} \right]$$


Since $K > \frac{L_F L_H}{4}$ holds true for this equilibrium, the denominator is positive. Hence, the first equilibrium is stable if the numerators of both eigenvalues have negative real parts, which turns out to be reduced to $K > \frac{L_F L_H}{2}$. The eigenvalues evaluated at the second equilibrium of (12) are

$$2K \left[ 4K - (L_F + L_H)^2 \pm \sqrt{(4K + (L_F + L_H)^2)^2 - 8L_F L_H (L_F + L_H)^2} \right]$$

$$L_F L_H (L_F + L_H)^2.$$

We can show that both eigenvalues have negative real parts if $K < \frac{L_F L_H}{2}$. Hence, we have shown that one of the two interior equilibria is always unstable, and thus, there exists a unique stable interior equilibrium, which is given by (12).

Next, we investigate the stability of corner equilibria. For small $K \leq \frac{(L_F + L_H) L_F}{4}$, there may exist corner equilibria. We have

$$\Delta V_H|_{(\lambda^*_F, \lambda^*_H) = (0, 1)} = 1 - \frac{4K}{L_H (L_F + L_H)} \geq 0,$$

$$\Delta V_F|_{(\lambda^*_F, \lambda^*_H) = (0, 1)} = 0,$$

implying that $(\lambda^*_F, \lambda^*_H) = (0, 1)$ is a corner equilibrium if $K \leq \frac{L_H (L_F + L_H)}{4}$. We also have

$$\Delta V_H|_{(\lambda^*_F, \lambda^*_H) = (1, 0)} = 0,$$

$$\Delta V_F|_{(\lambda^*_F, \lambda^*_H) = (1, 0)} = 1 - \frac{4K}{L_F (L_F + L_H)} \geq 0.$$

implying that $(\lambda^*_F, \lambda^*_H) = (1, 0)$ is a corner equilibrium if $K \leq \frac{L_F (L_F + L_H)}{4}$. Note that there are multiple stable corner equilibria if $K \leq \frac{L_F (L_F + L_H)}{4}$. Summarizing these results, we obtain Lemma 2.
Appendix B: Data by municipality

The number of foreign residents

The data on the number of residents by nationality and municipality in Tokyo prefecture is taken from the Tokyo Statistical Yearbook. It lists the top 18 countries for 1954-1968, top 7 nationalities for 1969-1982, top 8 nationalities for 1983-1998, and top 10 nationalities for 1999-2019. Therefore, the data exists for all years for major source countries like China and Korea, whereas the data for some years are missing for other source countries. The data covers all cities and wards in Tokyo prefecture, but the number of municipalities sometimes changes between 30 and 50 during the study period due to municipal merger and division.

The number of restaurants

In collecting data, we searched the number of restaurants on Tabelog by designating a municipality name as a location and entering a country name in a keyword field. For Filipino and Burmese restaurants, we first narrowed down candidates by choosing a category “All>Restaurants>Asian, Ethnic>Southeast Asian Cuisine>Southeast Asian Cuisine (and others)” and conducted the search. This is because, without narrowing down candidates, the search for Filipino restaurants mistakenly picks up users’ comments, such as “this Italian restaurant is located in front of the Philippine Embassy,” which results in a long list of unrelated restaurants. We collected the data on the number of Filipino restaurants on April 16, 2020, and that of the other four countries on April 11, 2020. Summary statistics of foreign cuisine restaurants are tabulated in Table A1.
Table A1: Summary statistics on the number of restaurants by cuisine

<table>
<thead>
<tr>
<th>Cuisine by country</th>
<th>Obs</th>
<th>Mean</th>
<th>Std.Dev.</th>
<th>Min</th>
<th>Max</th>
</tr>
</thead>
<tbody>
<tr>
<td>Chinese</td>
<td>49</td>
<td>180.3</td>
<td>170.4</td>
<td>18</td>
<td>680</td>
</tr>
<tr>
<td>Korean</td>
<td>49</td>
<td>42.8</td>
<td>69.5</td>
<td>1</td>
<td>430</td>
</tr>
<tr>
<td>Vietnamese</td>
<td>49</td>
<td>5.7</td>
<td>8.0</td>
<td>0</td>
<td>39</td>
</tr>
<tr>
<td>Filipino</td>
<td>49</td>
<td>0.4</td>
<td>0.6</td>
<td>0</td>
<td>2</td>
</tr>
<tr>
<td>Nepali</td>
<td>49</td>
<td>11.5</td>
<td>11.4</td>
<td>0</td>
<td>57</td>
</tr>
<tr>
<td>American</td>
<td>49</td>
<td>6.4</td>
<td>11.4</td>
<td>0</td>
<td>63</td>
</tr>
<tr>
<td>Indian</td>
<td>49</td>
<td>26.4</td>
<td>23.3</td>
<td>2</td>
<td>83</td>
</tr>
<tr>
<td>Burmese</td>
<td>49</td>
<td>0.5</td>
<td>2.0</td>
<td>0</td>
<td>10</td>
</tr>
<tr>
<td>Thai</td>
<td>49</td>
<td>16.3</td>
<td>20.5</td>
<td>0</td>
<td>84</td>
</tr>
</tbody>
</table>
Appendix C: Figures

Figure A1: Time series plots of $DI_i$ for other nationalities
Figure A2: Time series plots of $ID_i$
Figure A3: Time series plots of $CV_i$. 
Figure A4: The numbers of foreign residents and foreign restaurants by municipality, where the horizontal axis is $\frac{10^3L_{Fr}}{L_{Hr}}$ and the vertical axis is $\frac{10^4H_{Fr}}{L_{Hr}}$. 

(a) Vietnamese  
(b) Filipino  
(c) Nepali  
(d) Burmese  
(e) Indian  
(f) American  
(g) Thai
Figure A5: The adjusted numbers of foreign residents and foreign restaurants by municipality, where the horizontal axis is $\rho \sum_{s=1}^{3} \frac{P_{L_s}F_{s}}{PLH} + (1-\rho) \frac{P_{L_F}F_{F}}{PLH}$ and the vertical axis is $\rho \sum_{s=1}^{3} \frac{P_{N_s}F_{s}}{PLH} + (1-\rho) \frac{P_{N_F}F_{F}}{PLH}$. 